

Why Not Substitute $\frac{1}{2}$ DL for Nondetects?

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What's wrong with substitution?

- a. Produces **invasive data** alien to the concentrations actually in samples
- b. Adds a pattern to your data. Substitution is NOT neutral
- c. Produces poor estimates and incorrect statistical tests
- d. Far better methods are available. Right now. You don't need a PhD to do them.



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For More Information

on this topic and other methods for data analysis with nondetects, see our online course



Nondetects And Data Analysis
at <https://practicalstats.teachable.com>

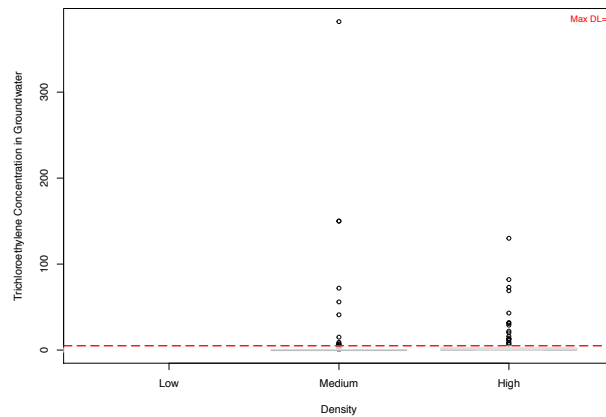


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Example: TCE concentrations in GW



- Reporting limits at 1, 2, 3, 4, and 5 ug/L. ~ 80% censoring
- Do TCE distributions differ among the three land-use groups?



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Substitution versus Kruskal-Wallis

ANOVA sub <5 = 0
p = 0.57

ANOVA sub <5 = 2.5
p = 0.56

ANOVA sub <5 = 5
p = 0.50

No differences found!

Kruskal-Wallis test
p = 0.01

No substitution

Distribution of TCE differs between land-use groups

Simplest Tests

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Better Tests For Censored Data

```

> cenlway (TCECONC, BDL_1, Density)
      n n.cen median  mean  sd
Factor=High  92  58 <5  7.778019 19.7826151
Factor=Low   25  23 <5  2.083333  0.5768489
Factor=Medium 130 113 <5  7.867264 39.7664396
                    
```

Oneway Peto-Peto test of CensData: TCECONC by Factor:
Density

Chisq = 16.25 on 2 degrees of freedom p = 0.000295

Distribution of TCE differs between land-use groups

Pairwise comparisons using Peto & Peto test

	High	Low
Low	0.0098	-
Medium	0.0024	0.2796

Shows which groups differ from others

Nonparametric Test using Multiple Detection Limits
from our Nondetects And Data Analysis course

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There is still much confusion over the value of nondetects

Dr. Ronald Hites in Environ. Science & Technology (2019) in a Scientific Opinion piece stated:

"... Any time you have a % non-detected >20%, for whatever reason, it is unlikely that the data set can give useful results."



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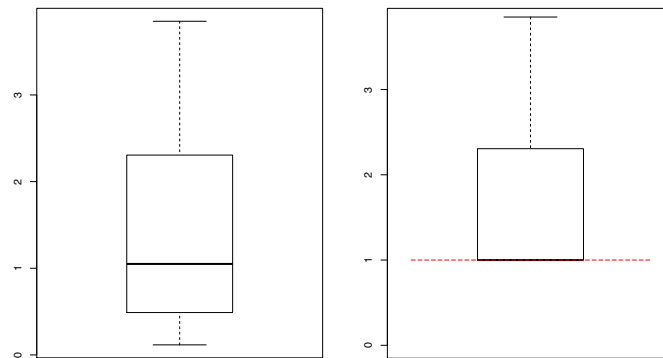
Substitution: Invasive data

Estimating the Standard Deviation

what happens to it when the same number is substituted for 60% of the observations?

=== All <2 become = 1 ===>

Before substituting.
True std. dev.
 $s = 1.02$



After substituting.
std. dev.
 $s = 0.84$

t-tests, ANOVA, regression, intervals all depend on the std. dev.



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Substitution when computing confidence intervals?

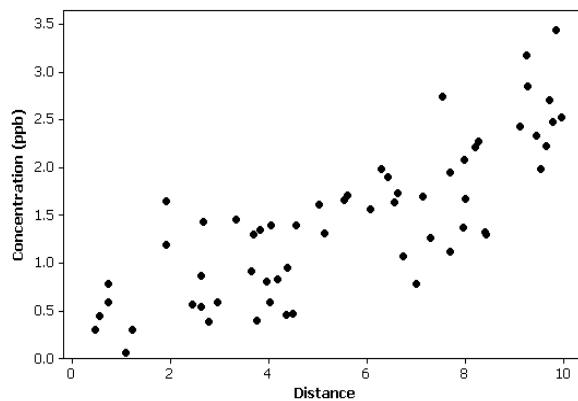
- Singh et al (2006), developers of the ProUCL software, determined that substituting $\frac{1}{2}$ DL “does not provide adequate coverage [UCL95 is not high enough] ...even for [% NDs] as low as 10%”
- Lower standard deviations produce lower confidence limits, too-short intervals.
- They summarize their results with "Do not use DL/2 (t) method to compute a UCL".
- Recommended methods were based on a Kaplan-Meier estimator, which is available in our course software.



Substitution: Invasive data

Correlation and Regression

Before censoring.
True correlation
 $r=0.81$



Substitution: Invasive data

Correlation and Regression

Two DLs @ 3 and 1.
 <3s become 1.5,
 <1s become 0.5.

After substitution.
 invasive data form flat
 (zero-slope) lines,
 lowering correlation
 to $r=0.55$

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Evaluation of Substitution for regression models

Thompson and Nelson (2003) found that for censored response (y) variables, substituting one-half the DL for nondetects produced

- biased parameter estimates (slopes) and
- artificially small standard error estimates (explanatory variables who shouldn't be in the regression appear to be significant)

There are better ways!

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Regression with censored data

No Substitution for Nondetects!

Software comes with NADA course

```
> Pbreg <- cencorreg(Blood, BloodCen, Kidney)
Likelihood R = 0.8236
Rescaled Likelihood R = 0.8721
McFaddens R = 0.714
> summary(Pbreg)
```

Correlation Coefficients

Regression Equation

Blood lead concentrations are related to lead in kidneys

	Value	Std. Error	z	p	
(Intercept)	-4.4573	0.1733	-25.72	< 2e-16	Loglik(model)= -14.7 Loglik(intercept only)= -30
Kidney	0.2436	0.0302	8.07	7.1e-16	Chisq= 30.62 on 1 degrees of freedom, p= 3.1e-08
Log(scale)	-0.6737	0.2036	-3.31	0.00094	

$$\ln(\text{blood Pb}) = -4.457 + 0.244 * \text{kidney Pb} \quad \text{or} \quad \text{blood Pb} = e^{-4.457} \cdot \text{kidneyPb}^{0.244}$$



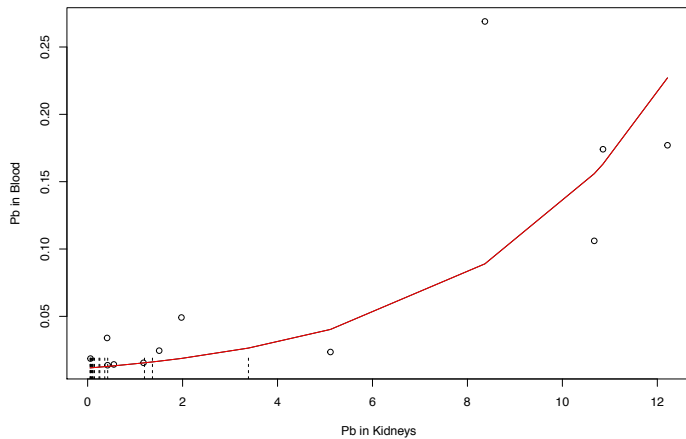
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Plotting the regression line

Regression straight line in log units becomes a curve in original units

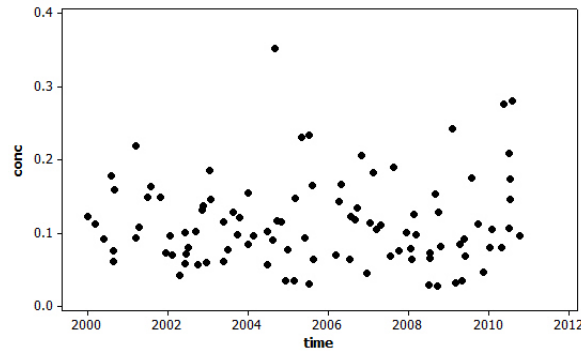
```
> cenxyplot(Kidney,
KidneyCen, Blood, BloodCen,
xlab = "Pb in Kidneys",
ylab = "Pb in Blood")
> ik <- order(Kidney)
> lines(Kidney[ik],
exp(predict(Pbreg)[ik]),
col = "red")
```

Nondetects are plotted as dashed vertical lines. The info that they occur only at the low kidney Pb concentrations is used in computing the regression model.



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Classic faux pas with substitution: finding a trend that isn't there



- No change over time. No trend.
- Will replace smallest values with a decreasing pattern of detection limits, mimicking what often happens in labs.

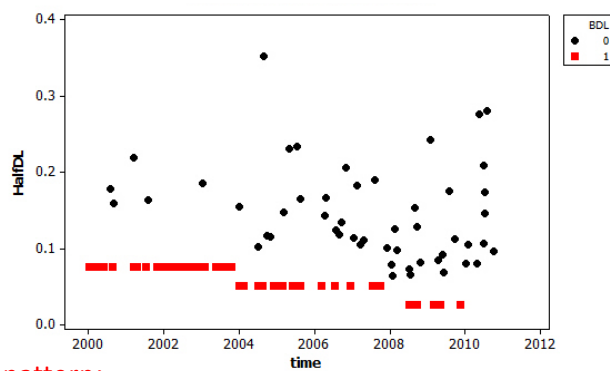


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Finding a trend that isn't there



Invasive pattern:

- DLs decrease over time.
- After substitution with $1/2DL$, decreases over time (trend) may test as significant



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Trend -- the censeaken function

The seasonal Kendall test for data with multiple DLs. No substitution. Available in our NADA course materials.

```
> censeaken (dectime, `Total Recoverable Chromium`, CrND, group = Season)
DATA ANALYZED: Total Recoverable Chromium vs dectime by Season
-----
Season N   S   tau   pval intercept   slope
1 Dry 34 -176 -0.314 0.0091337 79.103 -0.03901 Significant downtrend in Dry season
-----
Season N   S   tau   pval intercept   slope
1 Wet 29 -24 -0.0591 0.66604 24.355 -0.01169 No significant trend in Wet season
-----
Seasonal Kendall test and Theil-Sen line
N   S   Tau Pvalue_SK Nreps Intercept Slope
1 63 -200 -0.207 0.014 999 74.232 -0.03655 Significant trend overall. SK slope is -0.036 ug/L per year
-----
```

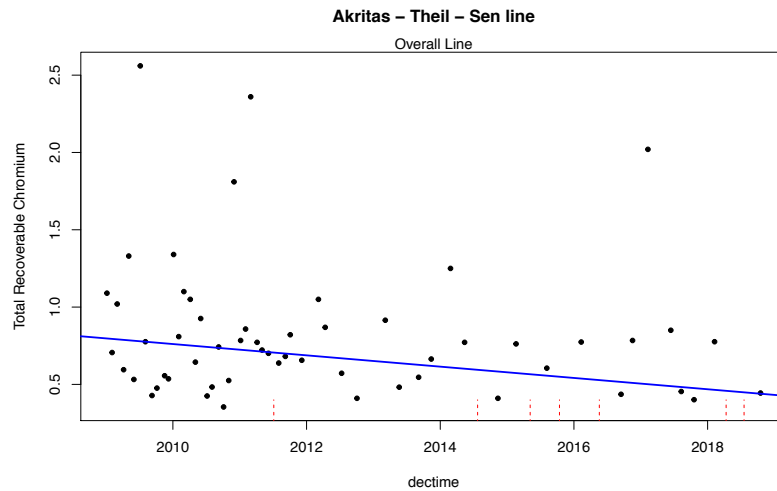


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Trend -- the censeaken function

Nondetects
(dashed vertical
lines) influence the
trend line and test.

They occur more
frequently at later
times, adding to
the evidence of a
downtrend.

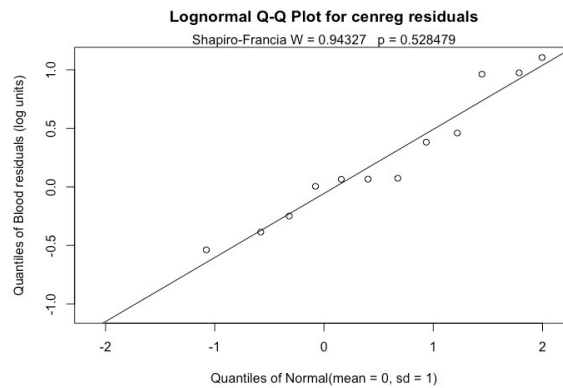


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Check Assumptions with QQ Plot

Regression is a parametric method:

Check the assumption of a normal distribution with a Q-Q plot of residuals. Here using the default of $\log(Y)$ fits well. It often does for data with NDs (because they are close to zero)



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What else is available?

- Our Nondetects And Data Analysis course covers:
 - parametric and nonparametric methods
 - computing summary statistics
 - confidence, prediction and tolerance intervals
 - plotting data
 - comparing to standards
 - testing differences of two groups
 - 3+ group tests with multiple comparisons
 - regression and correlation
 - trend analysis
 - multivariate methods

All without substitution of values like $1/2DL$ for nondetects

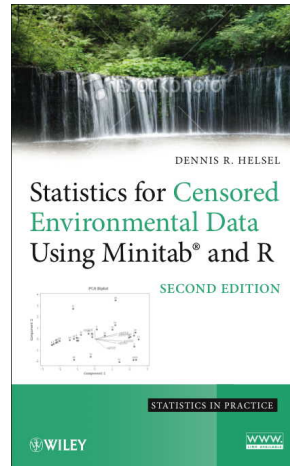


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For more detail:

Statistics for
Censored
Environmental
Data
(the second edition)

by Dennis R. Helsel
Wiley (2012)
www.PracticalStats.com/nada



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Free videos introducing methods for nondetects available now for streaming

on our Online Training Center
<https://practicalstats.teachable.com/>

Let colleagues know about them.

But they're only a fraction of the
material in the NADA course itself.

<1	<2	<3
Introduction to Nondetects And D... \$0 sales 276 enrolled	Fitting Distributions to Data with ... \$0 sales 68 enrolled	Testing Group Differences w/NDs \$0 sales 44 enrolled
<4	<5	<6
The Mystery of Nondetects \$0 sales 36 enrolled	Correlation and Regression for D... \$0 sales 18 enrolled	Trend Analysis for Data w/ NDs \$0 sales 5 enrolled



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Thank you for viewing

Statistical methods for data with nondetects at 1 or more detection limits is the focus of our Nondetects And Data Analysis course

Questions about this material or our courses?

Get in touch!

Dennis Helsel ask@practicalstats.com

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